# **Non-Bayesian Methods**

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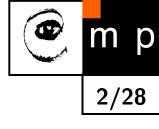
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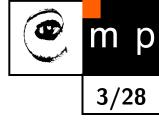
Oct 2023

# **Lecture Outline**

- 1. Limitations of Bayesian Decision Theory
- 2. Neyman Pearson Task
- 3. Minimax Task
- 4. Wald Task



# **Bayesian Decision Theory**



Recall:

- X set of observations
- K set of hidden states
- D set of decisions
- $p_{XK}$ :  $X \times K \rightarrow \mathbb{R}$ : joint probability
- $W: K \times D \rightarrow \mathbb{R}:$  loss function,
- $q: X \to D: \text{ strategy}$

R(q): risk:

$$R(q) = \sum_{x \in X} \sum_{k \in K} p_{XK}(x,k) \ W(k,q(x))$$
(1)

Bayesian strategy  $q^*$ :

$$q^* = \operatorname*{argmin}_{q \in X \to D} R(q) \tag{2}$$

# Limitations of the Bayesian Decision Theory



The limitations follow from the very ingredients of the Bayesian Decision Theory — the necessity to know all the probabilities and the loss function.

- The loss function W must make sense, but in many tasks it wouldn't
  - medical diagnosis task (W: price of medicines, staff labor, etc. but what penalty in case of patient's death?) Uncomparable penalties on different axes of X.
  - nuclear plant
  - judicial error
- The prior probabilities  $p_K(k)$ : must exist and be known. But in some cases it does not make sense to talk about probabilities because the events are not random.
  - $K = \{1, 2\} \equiv \{$ own army plane, enemy plane $\};$ p(x|1), p(x|2) do exist and can be estimated, but p(1) and p(2) don't.
- The conditionals may be subject to non-random intervention;  $p(x \mid k, z)$  where  $z \in Z = \{1, 2, 3\}$  are different interventions.
  - a system for handwriting recognition: The training set has been prepared by 3 different persons. But the test set has been constructed by one of the 3 persons only. This **cannot** be done:

(!) 
$$p(x \mid k) = \sum_{z} p(z)p(x \mid k, z)$$
 (3)

## **Neyman Pearson Task**

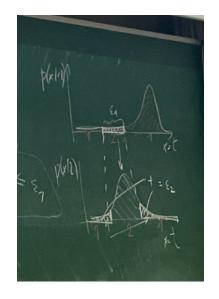


- X set of observations
- Conditionals p(x | 1), p(x | 2) are given
- p(KIA) The priors p(1) and p(2) are unknown or do not exist
- $q: X \to K$  strategy

The Neyman Pearson Task looks for the optimal strategy  $q^*$  for which

- i) the error of classification for class 1 is lower than a predefined threshold  $\bar{\epsilon}_1$  ( $0 < \bar{\epsilon}_1 < 1$ ), while
- ii) the classification error for class 2 is as low as possible.

This is formulated as an optimization task with an inequality constraint:



$$q^{*} = \underset{q:X \to K}{\operatorname{argmin}} \sum_{x:q(x) \neq 2} p(x \mid 2)$$
subject to: 
$$\sum_{x:q(x) \neq 1} p(x \mid 1) \leq \overline{\epsilon}_{1}.$$
(5)

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p(x/2)

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# Neyman Pearson Task



(copied from the previous slide:)

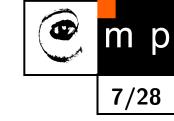
$$q^* = \underset{q:X \to K}{\operatorname{argmin}} \sum_{x:q(x)\neq 2} p(x \mid 2)$$
(4)  
subject to: 
$$\sum_{x:q(x)\neq 1} p(x \mid 1) \leq \overline{\epsilon}_1.$$
(5)

A strategy is characterized by the classification error values  $\epsilon_2$  and  $\epsilon_1$ :

$$\epsilon_{1} = \sum_{x:q(x)\neq 1} p(x \mid 1)$$

$$\epsilon_{2} = \sum_{x:q(x)\neq 2} p(x \mid 2)$$
(6)
(7)

# Example: Male/Female Recognition (Neyman Pearson) (1)



(8)

A hotel has an advertising screen in an elevator. Based on recognition of gender, it wants to display a relevant advert for a shopping mall located at the ground floor. The shopping mall is primarily designed to be interesting for female customers. For this reason, the female classification error threshold is set to  $\bar{\epsilon}_1 = 0.2$ . At the same time, the objective is to minimize mis-classification of male customers.

- $K = \{1, 2\} \equiv \{\mathsf{F}, \mathsf{M}\}$  (female, male)
- measurements X = height × weight (height sensor = simple optical sensor, weight sensor = standard component of elevators)
- height  $\in \{h_1, h_2, h_3\}$ , weight  $\in \{w_1, w_2, w_3, w_4\}$   $(h_1 < h_2 < h_3)$ ,  $(w_1 < w_2 < w_3 < w_4)$
- Prior probabilities do not exist.
- Conditionals are given as follows:

p(x F)					
$h_1$	.197	.145	.094	.017	
$h_2$	.077	.299	.145	.017	
$h_3$	.001	.008	.000	.000	
	$w_1$	$w_2$	$w_3$	$w_4$	

p(x M)					
$h_1$	.011	.005	.011	.011	
$h_2$	.005	.071	.408	.038	
$h_3$	.002	.014	.255	.169	
	$w_1$	$w_2$	$w_3$	$w_4$	

#### Neyman Pearson : Solution

The optimal strategy  $q^*$  for a given  $x \in X$  is constructed using the likelihood ratio  $\frac{p(x|2)}{p(x|1)}$ . Let there be a constant  $\mu \ge 0$ . Given this  $\mu$ , a strategy q is constructed as follows:

$$\frac{p(x \mid 2)}{p(x \mid 1)} > \mu \quad \Rightarrow \quad q(x) = 2,$$

$$\frac{p(x \mid 2)}{p(x \mid 1)} \le \mu \quad \Rightarrow \quad q(x) = 1.$$
(9)
(10)

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The optimal strategy  $q^*$  is obtained by selecting the minimal  $\mu$  for which there still holds that  $\epsilon_1 \leq \overline{\epsilon}_1$ .

Let us show this on an example.

# Example: Male/Female Recognition (Neyman Pearson) (2)



p(x 1)				
$h_1$	.197	.145	.094	.017
$h_2$	.077	.299	.145	.017
$h_3$	.001	.008	.000	.000
	$w_1$	$w_2$	$w_3$	$w_4$

	p(x  <b>2</b> )					
$h_1$	.011	.005	.011	.011		
$h_2$	.005	.071	.408	.038		
$h_3$	.002	.014	.255	.169		
	$  w_1$	$w_2$	$w_3$	$w_4$		

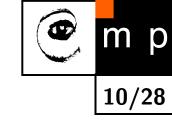
r(x) = p(x 2)/p(x 1)				
$h_1$	0.056	0.034	0.117	0.647
$h_2$	0.065	0.237	2.814	2.235
$h_3$	2.000	1.750	$\infty$	$\infty$
	$w_1$	$w_2$	$w_3$	$w_4$

rank order of $p(x 2)/p(x 1)$					
$h_1$	2	1	4	6	
$h_2$	3	5	10	9	
$h_3$	8	7	11	12	
	$w_1$	$w_2$	$w_3$	$w_4$	

Here, different  $\mu{\rm 's}$  can produce 11 different strategies.

First, let us take  $2.814 < \mu < \infty$ , e.g.  $\mu = 3$ . This produces a strategy  $q^*(x) = 1$  everywhere except where p(x|1) = 0. Obviously, classification error  $\epsilon_1 = 0$ , and  $\epsilon_2 = 1 - .255 - .169 = .576$ .

# Example: Male/Female Recognition (Neyman Pearson) (3)



p(x 1)					
$h_1$	.197	.145	.094	.017	
$h_2$	.077	.299	.145	.017	
$h_3$	.001	.008	.000	.000	
	$w_1$	$w_2$	$w_3$	$w_4$	

p(x  <b>2</b> )				
$h_1$	.011	.005	.011	.011
$h_2$	.005	.071	.408	.038
$h_3$	.002	.014	.255	.169
	$w_1$	$w_2$	$w_3$	$w_4$

r(x) = p(x 2)/p(x 1)				
$h_1$	0.056	0.034	0.117	0.647
$h_2$	0.065	0.237	2.814	2.235
$h_3$	2.000	1.750	$\infty$	$\infty$
	$w_1$	$w_2$	$w_3$	$w_4$

rank	rank, and $q^*(x) = \{1, 2\}$ for $\mu = 2.5$				
$h_1$	2	1	4	6	
$h_2$	3	5	10	9	
$h_3$	8	7	11	12	
	$w_1$	$w_2$	$w_3$	$w_4$	

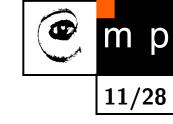
Next, take  $\mu$  which satisfies

$$r_9 < \mu < r_{10}$$
 (e.g.  $\mu = 2.5$ ) (11)

(where  $r_i$  is the likelihood ratios indexed by its rank.)

Here,  $\epsilon_1 = .145$ , and  $\epsilon_2 = 1 - .255 - .169 - .408 = .168$ .

# Example: Male/Female Recognition (Neyman Pearson) (4)



p(x 1)				
$h_1$	.197	.145	.094	.017
$h_2$	.077	.299	.145	.017
$h_3$	.001	.008	.000	.000
	$w_1$	$w_2$	$w_3$	$w_4$

p(x 2)				
$h_1$	.011	.005	.011	.011
$h_2$	.005	.071	.408	.038
$h_3$	.002	.014	.255	.169
	$w_1$	$w_2$	$w_3$	$w_4$

r(x) = p(x 2)/p(x 1)						
$h_1$	0.056	0.034	0.117	0.647		
$h_2$	0.065	0.237	2.814	2.235		
$h_3$	2.000	1.750	$\infty$	$\infty$		
	$w_1$	$w_2$	$w_3$	$w_4$		

rank	rank, and $q^*(x) = \{1, 2\}$ for $\mu = 2.1$					
$h_1$	2	1	4	6		
$h_2$	3	5	10	9		
$h_3$	8	7	11	12		
	$w_1$	$w_2$	$w_3$	$w_4$		

Do the same for  $\boldsymbol{\mu}$  satisfying

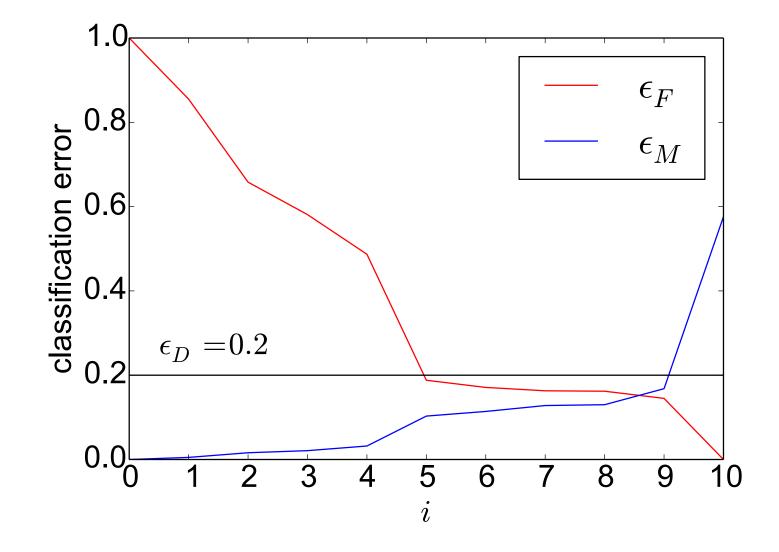
$$r_8 < \mu < r_9$$
 (e.g.  $\mu = 2.1$ ) (12)

 $\Rightarrow \epsilon_1 = .162$ , and  $\epsilon_2 = 0.13$ .

## Example: Male/Female Recognition (Neyman Pearson) (5)

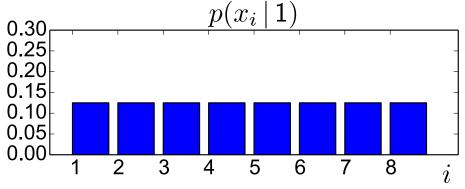


Classification errors for 1 and 2, for  $\mu_i = \frac{r_i + r_{i+1}}{2}$  and  $\mu_0 = 0$ .



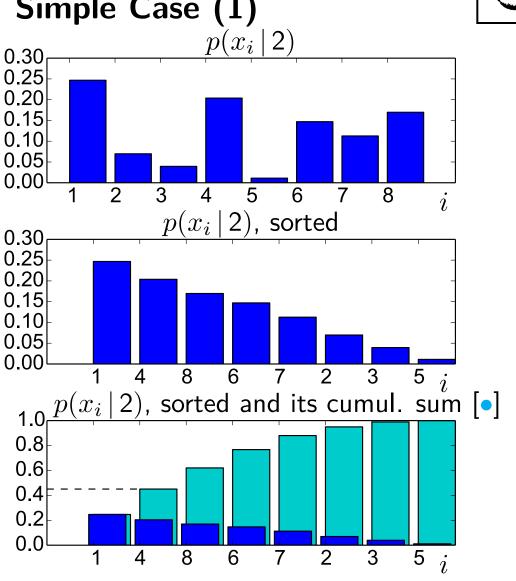
The optimum is reached for  $r_5 < \mu < r_6$ ;  $\epsilon_1 = .188$ ,  $\epsilon_2 = .103$ 

#### Neyman Pearson : Simple Case (1)



Consider a simple case when  $p(x_i | 1) = \text{const.}$  Possible values for  $\epsilon_1$ are  $0, \frac{1}{8}, \frac{2}{8}, ..., 1$ . If a strategy qclassifies P observations as normal then  $\epsilon_1 = \frac{P}{8}$ .

If P = 1 then  $\epsilon_1 = \frac{1}{8}$  and it is clear that  $\epsilon_2$  will attain minimum if the (one) observation which is classified as normal is the one with the highest  $p(x_i | 2)$ . Similarly, if P = 2 then the two observations to be classified as normal are the one with the first two highest  $p(x_i | 2)$ . Etc.



 $\uparrow$  cumulative sum of sorted  $p(x_i | 2)$  shows the classification success rate for 2, that is,  $1 - \epsilon_2$ , for  $\epsilon_1 = \frac{1}{8}, \frac{2}{8}, ..., 1$ . For example, for  $\epsilon_1 = \frac{2}{8}$   $(P = 2), \epsilon_2 = 1 - 0.45 = 0.55$ (as shown, dashed.)

# Neyman Pearson : Towards General Case (2)

In general,  $p(x_i | 1) \neq \text{const.}$  Consider the following example:

$p(x_i \mid 1)$			$p(x_i \mid 2)$			
$x_1$	$x_2$	$x_3$		$x_1$	$x_2$	$x_3$
0.5	0.25	0.25		0.6	0.35	0.05

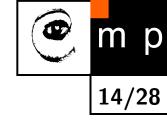
But this can easily be converted to the previous special case by (only formally) splitting  $x_1$  to two observations  $x'_1$  and  $x''_1$ :

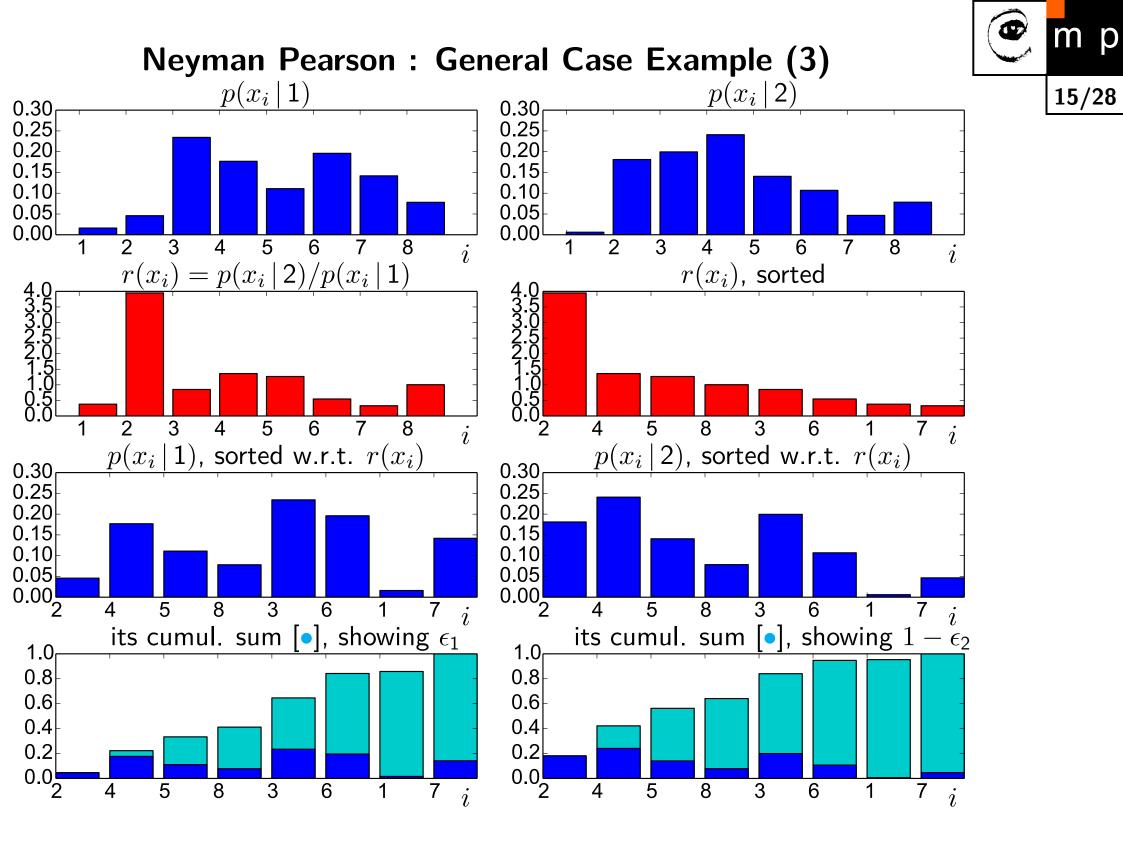
$\underline{\qquad} p(x_i   1)$				p(z)	$x_i   2)$			
$x'_1$	$x_1''$	$x_2$	$x_3$		$x'_1$	$x_1''$	$x_2$	$x_3$
0.25	0.25	0.25	0.25		0.3	0.3	0.35	0.05

which would result in ordering the observations by decreasing  $p(x_i | 2)$  as:  $x_2, x_1, x_3$ .

Obviously, the same ordering is obtained when  $p(x_i | 2)$  is 'normalized' by  $p(x_i | 1)$ , that is, using the likelihood ratio

$$r(x_i) = \frac{p(x_i \mid 2)}{p(x_i \mid 1)}.$$
(13)





#### Neyman Pearson Solution : Illustration of Principle

Lagrangian of the Neyman Pearson Task is

$$L(q) = \sum_{\substack{x: q(x)=1}} p(x \mid 2) + \mu \left( \sum_{x: q(x)=2} p(x \mid 1) - \bar{\epsilon}_D \right)$$
(14)  
=  $\overbrace{1 - \sum_{x: q(x)=2}}^{=} p(x \mid 2) + \mu \left( \sum_{x: q(x)=2} p(x \mid 1) \right) - \mu \bar{\epsilon}_1$ (15)  
=  $1 - \mu \bar{\epsilon}_1 + \sum_{x: q(x)=2} \underbrace{\{\mu p(x \mid 1) - p(x \mid 2)\}}_{T(x)}$ (16)

If T(x) is negative for an x then it will decrease the objective function and the optimal strategy  $q^*$  will decide  $q^*(x) = 2$ . This illustrates why the solution to the Neyman Pearson Task has the form

$$\frac{p(x \mid 2)}{p(x \mid 1)} > \mu \quad \Rightarrow \quad q(x) = 2,$$

$$\frac{p(x \mid 2)}{p(x \mid 1)} \le \mu \quad \Rightarrow \quad q(x) = 1.$$
(9)
(10)



#### Neyman Pearson : Derivation (1)



$$q^* = \min_{q:X \to K} \sum_{x:q(x)\neq 2} p(x \mid 2) \qquad \text{subject to:} \sum_{x:q(x)\neq 1} p(x \mid 1) \le \overline{\epsilon}_1.$$
(17)

Let us rewrite this as

$$q^* = \min_{q:X \to K} \sum_{x \in X} \alpha(x) p(x \mid 2) \qquad \text{subject to:} \qquad \sum_{x \in X} [1 - \alpha(x)] p(x \mid 1) \le \bar{\epsilon}_1. \tag{18}$$
  
and: 
$$\alpha(x) \in \{0, 1\} \ \forall x \in X \tag{19}$$

This is a combinatorial optimization problem. If the relaxation is done from  $\alpha(x) \in \{0, 1\}$  to  $0 \le \alpha(x) \le 1$ , this can be solved by **linear programming** (LP). The Lagrangian of this problem with inequality constraints is:

$$L(\alpha(x_1), \alpha(x_2), ..., \alpha(x_N)) = \sum_{x \in X} \alpha(x) p(x \mid 2) + \mu \left( \sum_{x \in X} [1 - \alpha(x)] p(x \mid 1) - \bar{\epsilon}_1 \right)$$
(20)  
$$- \sum_{x \in X} \mu_0(x) \alpha(x) + \sum_{x \in X} \mu_1(x) (\alpha(x) - 1)$$
(21)

# Neyman Pearson : Derivation (2)

$$L(\alpha(x_1), \alpha(x_2), ..., \alpha(x_N)) = \sum_{x \in X} \alpha(x) p(x \mid 2) + \mu \left( \sum_{x \in X} [1 - \alpha(x)] p(x \mid 1) - \bar{\epsilon}_1 \right)$$
(20)  
$$- \sum_{x \in X} \mu_0(x) \alpha(x) + \sum_{x \in X} \mu_1(x) (\alpha(x) - 1)$$
(21)

The conditions for optimality are  $(\forall x \in X)$ :

$$\frac{\partial L}{\partial \alpha(x)} = p(x \mid 2) - \mu p(x \mid 1) - \mu_0(x) + \mu_1(x) = 0, \quad (22)$$

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$$\mu \ge 0, \ \mu_0(x) \ge 0, \ \mu_1(x) \ge 0, \quad 0 \le \alpha(x) \le 1,$$
 (23)

$$\mu_0(x)\alpha(x) = 0, \ \mu_1(x)(\alpha(x) - 1) = 0, \ \mu\left(\sum_{x \in X} [1 - \alpha(x)]p(x \mid 1) - \bar{\epsilon}_1\right) = 0.$$
(24)

#### **Case-by-case analysis:**

case	implications
$\mu = 0$	L minimized by $\alpha(x) = 0  \forall x$
$\mu \neq 0,  \alpha(x) = 0$	$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x \mid 2) - \mu p(x \mid 1) \Rightarrow \frac{p(x \mid 2)}{p(x \mid 2)} \le \mu$
$\mu \neq 0,  \alpha(x) = 1$	$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x \mid 2) - \mu p(x \mid 1)] \Rightarrow \frac{p(x \mid 2)}{p(x \mid 2)} \ge \mu$
$egin{array}{ccc} \mu &  eq 0, \ 0 < lpha(x) < 1 \end{array}$	$\mu_0(x) = \mu_1(x) = 0 \Rightarrow \frac{p(x \mid 2)}{p(x \mid 1)} = \mu$

# Neyman Pearson : Derivation (3)



#### **Case-by-case analysis:**

case	implications
$\mu = 0$	L minimized by $\alpha(x) = 0  \forall x$
$\mu \neq 0,  \frac{\alpha(x) = 0}{\alpha(x)}$	$\mu_1(x) = 0 \Rightarrow \mu_0(x) = p(x \mid 2) - \mu p(x \mid 1) \Rightarrow \frac{p(x \mid 2)}{p(x \mid 2)} \le \mu$
$\mu \neq 0,  \alpha(x) = 1$	$\mu_0(x) = 0 \Rightarrow \mu_1(x) = -[p(x \mid 2) - \mu p(x \mid 1)] \Rightarrow \frac{p(x \mid 2)}{p(x \mid 2)} \ge \mu$
$\mu  eq 0, \ 0 < lpha(x) < 1$	$\mu_0(x) = \mu_1(x) = 0 \Rightarrow \frac{p(x \mid 2)}{p(x \mid 1)} = \mu$

**Optimal Strategy** for a given  $\mu \ge 0$  and particular  $x \in X$ :

$$\frac{p(x \mid 2)}{p(x \mid 1)} \quad \begin{cases} < \mu \quad \Rightarrow q(x) = 1 \text{ (as } \alpha(x) = 0) \\ > \mu \quad \Rightarrow q(x) = 2 \text{ (as } \alpha(x) = 1) \\ = \mu \quad \Rightarrow \text{LP relaxation does not give the desired solution, as } \alpha \notin \{0, 1\} \end{cases}$$
(25)



(26)

#### Consider:

p(x 1)				
$x_1$	$x_2$	$x_3$		
0.9	0.09	0.01		

	p(x 2)				
	$x_1$	$x_2$	$x_3$		
	0.09	0.9	0.01		

r(x) = p(x 2)/p(x 1)				
$x_1$	$x_2$	$x_3$		
0.1	10	1		

and  $\bar{\epsilon}_1 = 0.03$ .

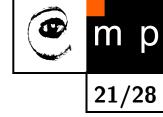
- $q_1: (x_1, x_2, x_3) \to (1, 1, 1) \Rightarrow \epsilon_1 = 0.00, \epsilon_2 = 1.00$
- $q_2: (x_1, x_2, x_3) \to (1, 1, 2) \implies \epsilon_1 = 0.01, \ \epsilon_2 = 0.99$
- ullet no other deterministic strategy q is feasible, that is all other ones have  $\epsilon_1 > ar \epsilon_1$
- q<sub>2</sub> is the best deterministic strategy but it does not comply with the previous basic result of constructing the optimal strategy because it decides for 2 for likelihood ratio 1 but decides for 1 for likelihood ratios 0.01 and 10. Why is that?
  - we can construct a randomized strategy which attains  $\overline{\epsilon}_1$  and reaches lower  $\epsilon_2$ :

$$q(x_1) = q(x_3) = 1, \quad q(x_2) = egin{cases} 2 & 1/3 \text{ of the time} \ 1 & 2/3 \text{ of the time} \ 1 & 2/3 \text{ of the time} \end{cases}$$

For such strategy,  $\epsilon_1 = 0.03$ ,  $\epsilon_2 = 0.7$ .

# Neyman Pearson : Note on Randomized Strategies (2)

- This is not a problem but a feature which is caused by discrete nature of X (does not happen when X is continuous).
- This is exactly what the case of  $\mu = p(x \mid 2)/p(x \mid 1)$  is on slide 18.

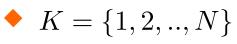


# Neyman Pearson : Notes (1)

- The task can be generalized to 3 hidden states, of which 2 are dangerous,  $K = \{2, D_1, D_2\}$ . It is formulated as an analogous problem with two inequality constraints and minimization of classification error for 2.
- Neyman's and Pearson's work dates to 1928 and 1933.
- A particular strength of the approach lies in that the likelihood ratio r(x) or even p(x | 2) need not be known. For the task to be solved, it is enough to know the p(x | 1) and the rank order of the likelihood ratio (to be demonstrated on the next page)

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## Minimax Task



- X set of observations
- Conditionals  $p(x \mid k)$  are known  $\forall k \in K$
- The priors p(k) are unknown or do not exist
- $q: X \to K$  strategy

The Minimax Task looks for the optimum strategy  $q^*$  which minimizes the classification error of the worst classified class:

$$q^* = \underset{q:X \to K}{\operatorname{argmin}} \max_{k \in K} \epsilon(k), \quad \text{where}$$

$$\epsilon(k) = \sum_{x:q(x) \neq k} p(x \mid k)$$
(27)
(28)

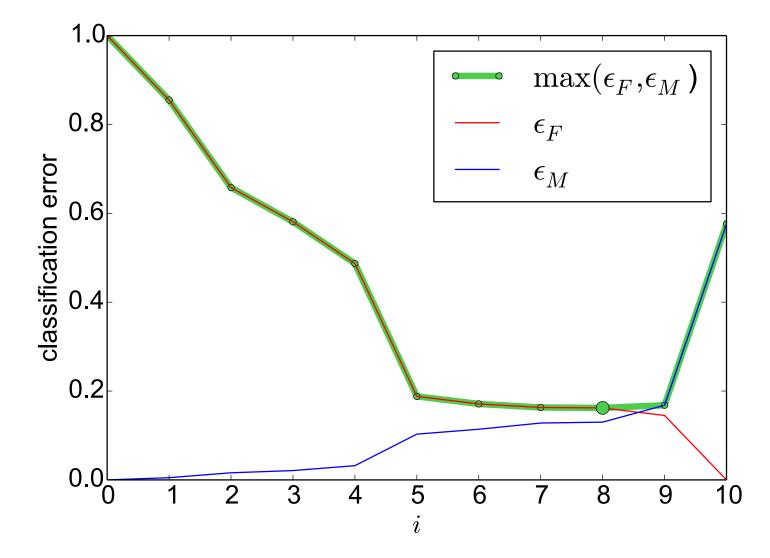
Example: A recognition algorithm qualifies for a competition using preliminary tests.
 During the final competition, only objects from the hardest-to-classify class are used.

- For a 2-class problem, the strategy is again constructed using the likelihood ratio.
- In the case of continuous observations space X, equality of classification errors is attained:  $\epsilon_1 = \epsilon_2$
- The derivation can again be done using Linear Programming.

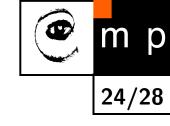


#### Example: Male/Female Recognition (Minimax)

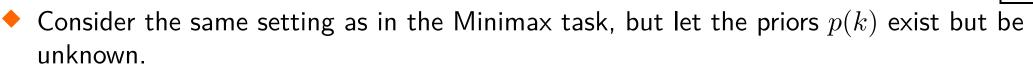
Classification errors for 1 and 2, for  $\mu_i = \frac{r_i + r_{i+1}}{2}$  and  $\mu_0 = 0$ .



The optimum is attained for i = 8,  $\epsilon_1 = .162$ ,  $\epsilon_2 = .13$ . The corresponding strategy is as shown on slide 11.



# Minimax: Comparison with Bayesian Decision with Unknown Priors



• The Bayesian error  $\epsilon$  for strategy q is

$$\epsilon = \sum_{k} \sum_{x: q(x) \neq k} p(x, k) = \sum_{k} p(k) \underbrace{\sum_{x: q(x) \neq k} p(x \mid k)}_{\epsilon(k)}$$
(29)

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- We want to minimize  $\epsilon$  but we do not know p(k)'s. What is the maximum it can attain? Obviously, the p(k)'s do the convex combination of the class errors  $\epsilon(k)$ ; the maximum Bayesian error will be attained when p(k) = 1 for the class k with the highest class error  $\epsilon(k)$ .
- Thus, to minimize the Bayesian error  $\epsilon$  under this setting, the solution is to minimize the error of the hardest-to-classify class.
- Therefore, Minimax formulation and the Bayesian formulation with Unknown Priors lead to the same solution.

# Wald Task (1)

- Let us consider classification with two states,  $K = \{1, 2\}$ .
- We want to set a threshold  $\epsilon$  on the classification error of both of the classes:  $\epsilon_1 \leq \epsilon$ ,  $\epsilon_2 \leq \epsilon$ .
- It is clear that there may be **no** feasible solution if  $\epsilon$  is set too low.
- That is why the possibility of decision "do not know" is introduced. Thus  $D = K \cup \{?\}$

• A strategy  $q: X \to D$  is characterized by:

$$\epsilon_1 = \sum_{x: q(x)=2} p(x \mid 1) \quad \text{(classification error for 1)} \tag{30}$$

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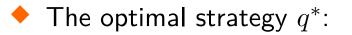
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$$\epsilon_2 = \sum_{x: q(x)=1} p(x \mid 2) \quad \text{(classification error for 2)} \tag{31}$$

$$\kappa_1 = \sum_{x: q(x)=?} p(x \mid 1) \quad \text{(undecided rate for 1)}$$
(32)

$$\kappa_2 = \sum_{x: q(x)=?} p(x \mid 2) \quad \text{(undecided rate for 2)}$$
(33)

# Wald Task (2)



$$q^* = \underset{q:X \to D}{\operatorname{argmin}} \max_{i=\{1,2\}} \kappa_i$$
subject to:  $\epsilon_1 \le \epsilon, \epsilon_2 \le \epsilon$ 
(34)
(35)

The task is again solvable using LP (even for more than 2 classes)

The optimal solution is again based on the likelihood ratio

$$r(x) = \frac{p(x \mid 1)}{p(x \mid 2)}$$
(36)

The optimal strategy is constructed using suitably chosen thresholds  $\mu_l$  and  $\mu_h$  such that:

$$q(x) = \begin{cases} 2 & \text{for } r(x) < \mu_l \\ 1 & \text{for } r(x) > \mu_h \\ ? & \text{for } \mu_l \le r(x) \le \mu_h \end{cases}$$
(37)





# Example: Male/Female Recognition (Wald)

Solve the Wald task for  $\epsilon=0.05.$ 

		p(x F)	)			
$h_1$	.197	.145	.094	.017	$h_1$	.01
$h_2$	.077	.299	.145	.017	$h_2$	.00
$h_3$	.001	.008	.000	.000	$h_3$	.00
	$w_1$	$w_2$	$w_3$	$w_4$		$\mid w$

p(x M)						
$h_1$	.011	.005	.011	.011		
$h_2$	.005	.071	.408	.038		
$h_3$	.002	.014	.255	.169		
	$  w_1$	$w_2$	$w_3$	$w_4$		

	r(x) = p(x 2)/p(x 1)						
$h_1$	0.056	0.034	0.117	0.647			
$h_2$	0.065	0.237	2.814	2.235			
$h_3$	2.000	1.750	$\infty$	$\infty$			
	$w_1$	$w_2$	$w_3$	$w_4$			

rank, and $q^*(x) = \{1, 2, ?\}$						
-7		$h_1$	2	1	4	6
5		$h_2$	3	5	10	9
		$h_3$	8	7	11	12
1			$w_1$	$w_2$	$w_3$	$w_4$

**Result:**  $\epsilon_2 = 0.032$ ,  $\epsilon_1 = 0$ ,  $\kappa_2 = 0.544$ ,  $\kappa_1 = 0.487$ 

$$(r_4 < \mu_l < r_5, r_{10} < \mu_h < \infty)$$